# Homework 6 Math 202 Stochastic Processes Spring 2024 

Question 1. Consider the Markov chain described in HW3 Q2.
(a) After a long time, what would be the expected number of papers in the pile?
(b) Assume the pile starts with 0 papers. What is the expected time until the pile will again have 0 papers?

## Solution:

(a) Recall from first homework that this is a Markov chain with state space $\{0,1,2,3,4\}$ and transition matrix
$\left[\begin{array}{ccccc}\frac{1}{3} & \frac{2}{3} & 0 & 0 & 0 \\ \frac{1}{3} & 0 & \frac{2}{3} & 0 & 0 \\ \frac{1}{3} & 0 & 0 & \frac{2}{3} & 0 \\ \frac{1}{3} & 0 & 0 & 0 & \frac{2}{3} \\ 1 & 0 & 0 & 0 & 0\end{array}\right]$.

Let us start finding stationary distribution. Solving $\pi P=\pi$ for a probability vector, we get

$$
\pi=\left(\frac{81}{211}, \frac{54}{211}, \frac{36}{211}, \frac{24}{211}, \frac{16}{211}\right)
$$

Hence expected number of papers in the pile in the long run turns out to be

$$
0 \times \frac{81}{211}+1 \times \frac{54}{211}+2 \times \frac{36}{211}+3 \times \frac{24}{211}+4 \times \frac{16}{211}=\frac{262}{211} \approx 1.24
$$

(b) $E[T]=\frac{1}{\pi(0)}=\frac{211}{81} \approx 2.6$.

Question 2. Suppose we flip a fair coin repeatedly until we have flipped four consecutive heads. What is the expected number of flips that are needed? (Hint: consider a Markov chain with state space $\{0,1,2,3,4\}$.)

## Solution:

These probabilities can be modelled with a markov chain with a state space $\{0,1,2,3,4\}$ where $p(i, 0)=$ $1 / 2$ where we flipped a tail and $p(i, i+1)=1 / 2$ if we flip head in the next flip for $i=0,1,2,3$ and $p(4,4)=1$ because 4 is absorbing. The transition matrix corresponding to this chain becomes

$$
P=\left[\begin{array}{ccccc}
1 / 2 & 1 / 2 & 0 & 0 & 0 \\
1 / 2 & 0 & 1 / 2 & 0 & 0 \\
1 / 2 & 0 & 0 & 1 / 2 & 0 \\
1 / 2 & 0 & 0 & 0 & 1 / 2 \\
0 & 0 & 0 & 0 & 1
\end{array}\right] .
$$

We will repeat the same computations.

$$
\left.\tilde{P}=\begin{array}{c}
4 \\
0
\end{array} \begin{array}{r}
4 \\
0 \\
1 \\
2 \\
2
\end{array} \begin{array}{rrrrr}
1 & 0 & 0 & 0 & 0 \\
0 & 1 / 2 & 1 / 2 & 0 & 0 \\
0 & 1 / 2 & 0 & 1 / 2 & 0 \\
0 & 1 / 2 & 0 & 0 & 1 / 2 \\
1 / 2 & 1 / 2 & 0 & 0 & 0
\end{array}\right] .
$$

$$
\begin{aligned}
M=(I-Q)^{-1} & =\left[\begin{array}{cccc}
1 / 2 & -1 / 2 & 0 & 0 \\
-1 / 2 & 1 & -1 / 2 & 0 \\
-1 / 2 & 0 & 1 & -1 / 2 \\
-1 / 2 & 0 & 0 & 1
\end{array}\right]^{-1} \\
& =\left[\begin{array}{llll}
16 & 8 & 4 & 2 \\
14 & 8 & 4 & 2 \\
12 & 6 & 4 & 2 \\
8 & 4 & 2 & 2
\end{array}\right]
\end{aligned}
$$

Since we don't have any tails in the beginning expected number of steps needed turns out to be $16+8+$ $4+2=30$.

Question 3. Consider the Markov chain with state space $S=\{0,1,2 \cdots\}$ and transition probabilities:

$$
p(x, x+1)=2 / 3 ; \quad p(x, 0)=1 / 3
$$

Show that the chain is positive recurrent and give the limiting probability $\pi$.

## Solution:

Writing the equations corresponding to $\pi=\pi P$ where $\sum_{x \in S} \pi(x)=1$, we get

$$
\begin{align*}
& \pi(0)=\frac{1}{3} \sum_{x \in S} \pi(x)=\frac{1}{3}  \tag{1}\\
& \pi(x)=\frac{2}{3} \pi(x-1) . \tag{2}
\end{align*}
$$

Hence the chain is positive recurrent with the limiting distribution $\pi(x)=\frac{1}{3}\left(\frac{2}{3}\right)^{x}, x \in S$.

Question 4. Consider the simple $(p=q=1 / 2)$ random walk on integers $\mathbb{Z}$. We argued in class that

$$
p_{n}(0,0) \sim \frac{C}{\sqrt{n}}
$$

for some constant $C$.
(a) Show that any two states $i$ and $j$ communicate.
(b) Show that 0 is recurrent by showing that the sum $\sum_{n} p_{n}(0,0)$ is divergent (consider the integral test for series)
(c) Is the symmetric random walk null recurrent or positive recurrent?

## Solution:

(a) Given any state $i<j$. Let $n=|i-j|$. Then

$$
\begin{align*}
& p_{n}(i, j)=\left(\frac{1}{2}\right)^{n}>0  \tag{3}\\
& p_{n}(j, i)=\left(\frac{1}{2}\right)^{n}>0
\end{align*}
$$

(b)

$$
\begin{equation*}
\sum_{n=1}^{\infty} \frac{1}{\sqrt{n}} \sim \int_{1}^{\infty} \frac{1}{\sqrt{x}} d x=\infty \tag{4}
\end{equation*}
$$

It implies that

$$
\begin{equation*}
\sum_{n=1}^{\infty} p_{n}(0,0)=\infty \tag{5}
\end{equation*}
$$

and thus 0 is recurrent.By part a, the chain is irreducible, so the chain is null recurrent.
(c) Symmetric random walk on $\mathbb{Z}$ is null recurrent since

$$
\begin{equation*}
\lim _{n \rightarrow \infty} p_{n}(0,0)=0 \tag{6}
\end{equation*}
$$

Hence, 0 is null recurrent. By part a, the chain is irreducible, so the chain is null recurrent.

Question 5. Consider the Markov chain with state space $S=\{0,1,2 \cdots\}$ and transition probabilities:

$$
\begin{array}{r}
p(x, x+2)=p ; \quad p(x, x-1)=1-p \quad x>0 \\
p(0,2)=p, \quad p(0,0)=1-p
\end{array}
$$

For which values of $p$ is this a transient chain? Hint: Use the Stirling's approximation:

$$
n!\sim \sqrt{2 \pi n}\left(\frac{n}{e}\right)^{n}
$$



## Solution:

Note that this chain has period 3 and $p_{3 n}(x, x)=\binom{3 n}{n} p^{n}(1-p)^{2 n}$. Then the chain is transient if $\sum_{n=1}^{\infty} p_{3 n}(x, x)<\infty$. To see for which $p$ 's this series converges, let us use Stirling's approximation:

$$
p_{3 n}(x, x) \approx \frac{\sqrt{6 \pi n}(3 n)^{3 n} e^{-3 n}}{\sqrt{4 \pi n}(2 n)^{2 n} e^{-2 n} \sqrt{2 \pi n} n^{n} e^{-n}}\left(p(1-p)^{2}\right)^{n}=\frac{C}{\sqrt{n}}\left(\frac{27}{4} p(1-p)^{2}\right)^{n}
$$

So, $\sum_{n=0}^{\infty} p_{3 n}(x, x)<\infty$ only if $\frac{27}{4} p(1-p)^{2}<1$ only if $p \in[0,1] \backslash\left\{\frac{1}{3}\right\}$.

Question 6. Consider the queueing model. For which values of $p, q$ is the chain null recurrent, positive recurrent, transient? For the positive recurrent case give the limiting probability distribution $\pi$. What is the average length of the queue in equilibrium? For the transient case, give $\alpha(x)=$ the probability starting at $x$ of ever reaching at 0 .

## Solution:

This is a Markov chain with state space $S=\{0,1,2,3 \cdots\}$ and transition probabilities

$$
\left.\left.\begin{array}{rlrl}
p(x, x-1)=q(1-p), & p(x, x) & =q p+(1-q)(1-p), & p(x, x+1)
\end{array}\right)=p(1-q), \quad x>0\right)
$$

Let $z \in S$ be fixed. Recall that the chain is transient if there exists a unique function $\alpha(x), x \in S$ satisfying

1. $\alpha(z)=1$ and $0 \leq \alpha(x) \leq 1$ for all $x \in S$.
2. $\alpha(x)=\sum_{y \in S} p(x, y) \alpha(y)$ and
3. $\inf \{\alpha(x): x \in S\}=0$.

Let $z=0$. Then $\alpha(0)=1$ and using (2) $\alpha$ satisfies for $x \geq 1$ :

$$
\begin{equation*}
\alpha(x)=q(1-p) \alpha(x-1)+(q p+(1-p)(1-q)) \alpha(x)+p(1-q) \alpha(x+1) \tag{7}
\end{equation*}
$$

Looking for a solution of the form $\alpha(x)=c^{x}$ for some constant $c$ we obtain c satisfies the quadratic equation

$$
p(1-q) c^{2}+(2 p q-p-q) c+q(1-p)=0
$$

The roots are

$$
\begin{equation*}
c_{1,2}=\frac{p+q-2 p q \pm|p-q|}{2 p(1-q)} \tag{8}
\end{equation*}
$$

and hence the general solution to (1) becomes

$$
\alpha(x)= \begin{cases}a_{1}+a_{2}\left(\frac{q(1-p)}{p(1-q)}\right)^{x} & p>q  \tag{9}\\ a_{1}+a_{2} x & p=q \\ a_{1}+a_{2}\left(\frac{q(1-p)}{p(1-q)}\right)^{x} & q>p\end{cases}
$$

Using $\alpha(0)=1$, we obtain

$$
\alpha(x)= \begin{cases}a+(1-a)\left(\frac{q(1-p)}{p(1-q)}\right)^{x} & p>q  \tag{10}\\ 1+a x & p=q \\ a+(1-a)\left(\frac{q(1-p)}{p(1-q)}\right)^{x} & q>p\end{cases}
$$

In the cases $p=q$ and $q>p$ the conditions $0 \leq \alpha(x) \leq 1$ and $\inf \{\alpha(x): x \in S\}=0$ cannot hold simultaneously since $x$ and $\left(\frac{q(1-p)}{p(1-q)}\right)^{x}$ are increasing. In the case $p>q$, $\inf \{\alpha(x): x \in S\}=0$ implies $a=0$. So the Markov chain is transient if $p>q$ and by uniqueness

$$
\alpha(x):=P\left(X_{n}=0 \text { for some } n \geq 0 \mid X_{0}=x\right)=\left(\frac{q(1-p)}{p(1-q)}\right)^{x}
$$

To decide between null recurrent and positive recurrent, let us see in which case we can find a nontrivial stationary distribution $\pi$. Solving for $\pi=\pi P$, we get

$$
\begin{array}{r}
\pi(x)=p(x-1, x) \pi(x-1)+p(x, x) \pi(x)+p(x+1, x) \pi(x+1) \\
\pi(0)=p(0,0) \pi(0)+p(1,0) \pi(1) .
\end{array}
$$

Using the probabilities, we have

$$
\begin{array}{r}
\pi(x)=p(1-q) \pi(x-1)+(q p+(1-q)(1-p)) \pi(x)+q(1-p) \pi(x+1) \\
\pi(0)=(1-p) \pi(0)+q(1-p) \pi(1) .
\end{array}
$$

Using the same method as above we get $\pi(0)=\frac{q(1-p)}{p} \pi(1)=c_{1}(1-q)$ and for $x \geq 1$

$$
\pi(x)= \begin{cases}a_{1}+a_{2}\left(\frac{p(1-q)}{q(1-p)}\right)^{x} & q>p  \tag{11}\\ a_{1}+a_{2} x & p=q\end{cases}
$$

Since $\pi$ is a probability vector $a_{1}=0$ and in the case $p=q$ no such $\pi$ exists. So the case $p=q$ is null recurrent. Now let's choose $a_{2}$ so that $\pi$ is a probability distribution:

$$
\sum_{x \in S} \pi(x)=a_{2}\left(1-q+\sum_{x=1}^{\infty}\left(\frac{p(1-q)}{q(1-p)}\right)^{x}\right)=a_{2}\left(1-q+\frac{p(1-q)}{q-p}\right)=a_{2} \frac{q(1-q)}{q-p}=1
$$

so $a_{2}=\frac{q-p}{q(1-q)}$ and for $x \geq 1$ :

$$
\begin{equation*}
\pi(x)=\frac{q-p}{q(1-q)}\left(\frac{p(1-q)}{q(1-p)}\right)^{x} \tag{12}
\end{equation*}
$$

and $\pi(0)=\frac{q-p}{q}$ and in the case $q>p$, the Markov chain is positive recurrent.

